Discovering Shape Categories by Clustering Shock Trees

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Abstract. This paper investigates whether meaningful shape categories can be identified in an unsupervised way by clustering shock-trees. We commence by computing weighted and unweighted edit distances between shock-trees extracted from the Hamilton-Jacobi skeleton of 2D binary shapes. Next we use an EM-like algorithm to locate pairwise clusters in the pattern of edit-distances. We show that when the tree edit distance is weighted using the geometry of the skeleton, then the clustering method returns meaningful shape categories.

Keywords: clustering shock trees, EM algorithm, Hamilton-Jacobi skeleton

1 Introduction

There has recently been considerable interest in the use of the reaction-diffusion equation as a means of representing and analysing both 2D and 3D shapes [1, 2,3]. In a nutshell, the idea is to extract a skeletal representation by evolving the shape-boundary inwards until singularities appear. Through the analysis of the differential properities of the singularities, a structural abstraction of the skeleton known as the shock-graph may be extracted. Although this abstraction has been widely used for shape-matching and recognition [2], its use as a means of learning shape categories has attracted less attention. The aim in this paper is to investigate whether graph-clustering can be used as a means of partitioning shock-trees into shape classes via unsupervised learning.

Graph clustering is an important yet relatively under-researched topic in machine learning [4,5]. The importance of the topic stems from the fact that it is an important tool for learning the class-structure of data abstracted in terms of relational graphs. Problems of this sort are posed by a multitude of unsupervised learning tasks in knowledge engineering, pattern recognition and computer vision. The process can be used to structure large data-bases of relational models [6] or to learn equivalence classes. One of the reasons for limited progress in the area has been the lack of algorithms suitable for clustering relational structures. In particular, the problem has proved elusive to conventional central clustering techniques. The reason for this is that it has proved difficult to define what is meant by the mean or representative graph for each cluster. However, Munger,

Bunke and Jiang [7] have recently taken some important steps in this direction by developing a genetic algorithm for searching for median graphs. A more fruitful avenue of investigation may be to pose the problem as pairwise clustering. This requires only that a set of pairwise distances between graphs be supplied. The clusters are located by identifying sets of graphs that have strong mutual pairwise affinities. There is therefore no need to explicitly identify an representative (mean, mode or median) graph for each cluster. Unfortunately, the literature on pairwise clustering is much less developed than that on central clustering.

When posed in a pairwise setting, the graph-clustering problem requires two computational ingredients. The first of these is a distance measure between relational structures. The second is a means of performing pairwise clustering on the distance measure. There are several distance measures available in the literature. For instance, in the classical pattern recognition literature, Haralick and Shapiro [8] have described a relational distance measure between structural descriptions. while Eshera and Fu [9] have extended the concept of edit distance from strings to graphs. More recently, Christmas, Kittler and Petrou [10], Wilson and Hancock [11] and Huet and Hancock [12] have developed probabilistic measures of graph-similarity. Turning our attention to pairwise clustering, there are several possible routes available. The simplest is to transform the problem into a central clustering problem. For instance, it is possible to embed the set of pairwise distances in a Euclidean space using a technique such as multi-dimensional scaling and to apply central clustering to the resulting embedding. The second approach is to use a graph-based method [13] to induce a classification tree on the data. Finally, there are mean-field methods which can be used to iteratively compute cluster-membership weights [14]. These methods require that the number of pairwise clusters be known a priori.

Our appraoch is as follows. Commencing from a data-base of silhouettes, we extract the Hamilton-Jacobi skeleton and locate the shocks which correspond to singularities in the evolution of the object boundary under the eikonal equation. We compute the similarity of the shapes using weighted and un-weighted tree edit distance. With the set of pairwise edit-distances between the shockgraphs to hand, we use a maximum-likelihood method for pairwise clustering. Our experiments show that when used in conjunction with the weighted tree edit distance, the pairwise clustering process locates meaningful shape categories.

2 Shock Tree Edit Distance

The practical problem tackled in this paper is the clustering of 2D binary shapes based on the similarity of their shock-trees. The idea of characterizing boundary shape using the differential singularities of the reaction equation was first introduced into the computer vision literature by Kimia, Tannenbaum and Zucker [3]. The idea is to evolve the boundary of an object to a canonical skeletal form using the reaction-diffusion equation. The skeleton represents the singularities in the curve evolution, where inward moving boundaries collide. The reaction component of the boundary motion corresponds to morphological erosion of the

| | 1 | 1 | 1 | 1 | 1 | 1 | 4 | 1 | 1 | 4 | - | d | - | 1 | 4 | 1 |
|----|---|--|---------------------------|-------|-------|-------|-------|-------|-------|-------|------------------------------|-------|-------|-------|-------|-------|
| 1 | 0.981 | 0.844 | 0.434 | 0.604 | 0.562 | 0.600 | 0 497 | 0.554 | 0.422 | 0.511 | 0.555 | | 14 | | | 0.402 |
| 1 | | | | | | | | | | | | | | | | 0.402 |
| ` | | | | | | | | | 1 | | | | | | | 0.434 |
| 1 | | | | | | | | | | | | | | | | 0.404 |
| 1 | | | | | | | | | | | | | | | | 0.404 |
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| _ | | | | | | | | | | | | | | | 0.353 | |
| 4 | Maria de la compansión | Commence of the Commence of th | The state of the state of | | 100 | | | | | | A CONTRACTOR OF THE PARTY OF | | - | _ | 0.630 | _ |
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| ef | | | | | | | | | | | | | | | 0.519 | |
| 14 | | | | | | | | | | | | | | | 0.687 | |
| * | 0.443 | 0.520 | 0.334 | 0.436 | 0.570 | 0.536 | 0.465 | 0.422 | 0.335 | 0.632 | 0.554 | 0.494 | 0.696 | 0.976 | 0.895 | 0.840 |
| 4 | 0.431 | 0.450 | 0.409 | 0.425 | 0.520 | 0.520 | 0.480 | 0.445 | 0.378 | 0.629 | 0.543 | 0.572 | 0.719 | 0.847 | 0.982 | 0.771 |
| | 0.441 | 0.447 | 0.380 | 0.447 | 0.560 | 0.520 | 0.493 | 0.434 | 0.346 | 0.563 | 0.558 | 0.541 | 0.765 | 0.864 | 0.817 | 1.000 |

Fig. 1. Pairwise edit distances computed using un-weighted trees.

boundary, while the diffusion component introduces curvature dependent boundary smoothing. Once the skeleton is to hand, the next step is to devise ways of using it to characterize the shape of the original boundary. Here we follow Zucker, Siddiqi, and others, by labeling points on the skeleton using so-called shock-classes [2]. We abstract the skeletons as trees in which the level in the tree is determined by their time of formation [15,2]. The later the time of formation, and hence their proximity to the center of the shape, the higher the shock in the hierarchy. While this temporal notion of relevance can work well with isolated shocks (maxima and minima of the radius function), it fails on monotonically increasing or decreasing shock groups. To give an example, a protrusion that ends on a vertex will always have the earliest time of creation, regardless of its relative relevance to the shape.

To overcome this drawback, we augment the structural information given by the skeleton topology and the relative time of shock formation, with a measure of feature importance. We opt to use a shape-measure based on the rate of change of boundary length with distance along the skeleton. To compute the measure we construct the osculating circle to the two nearest boundary points at each location on the skeleton.

This measurement has previously been used in the literature to express *relevance* of a branch when extracting or pruning the skeleton, but is has recently been shown that its geometric and differential properties make it a good measure of shape similarity [16].

| | 1 | / | - | 1 | 1 | 1 | 4 | 1 | 1 | 4 | - | et | A | * | * | * |
|----|-------|---------|-----------|------------|----------|-------|-------|-------|-----------|----------|------------|-------|-----------|-------|-------|-------------|
| 1 | 1.000 | 1.000 | 0.774 | 1.000 | 0.889 | 0.889 | 0.706 | 0.643 | 0.850 | 0.818 | 0.880 | 0.635 | 0.640 | 0.706 | 0.604 | 0.694 |
| / | | 1157130 | No. | | \$1150 B | 0.889 | 1000 | | | | | | | | 7 | |
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| 1 | 0.889 | 0.889 | 0.833 | 0.889 | 1.000 | 1.000 | 0.765 | 0.730 | 0.950 | 0.808 | 0.778 | 0.673 | 0.680 | 0.765 | 0.750 | 0.737 |
| 1 | 0.706 | 0.706 | 0.676 | 0.706 | 0.765 | 0.765 | 1.000 | 0.782 | 0.794 | 0.674 | 0.680 | 0.730 | 0.692 | 0.765 | 0.801 | 0.669 |
| 1 | 0.643 | 0.643 | 0.714 | 0.643 | 0.730 | 0.730 | 0.847 | 1.000 | 0.771 | 0.649 | 0.639 | 0.769 | 0.724 | 0.651 | 0.698 | 0.682 |
| / | 0.850 | 0.850 | 0.800 | 0.850 | 0.950 | 0.950 | 0.794 | 0.857 | 1.000 | 0.764 | 0.739 | 0.692 | 0.700 | 0.794 | 0.778 | 0.763 |
| * | 0.818 | 0.818 | 0.773 | 0.818 | 0.808 | 0.808 | 0.599 | 0.731 | 0.764 | 1.000 | 0.909 | 0.647 | 0.720 | 0.749 | 0.806 | 0.789 |
| * | 0.889 | 0.889 | 0.694 | 0.889 | 0.778 | 0.778 | 0.680 | 0.548 | 0.739 | 0.909 | 0.778 | 0.598 | 0.680 | 0.595 | 0.667 | 0.737 |
| d | 0.635 | 0.635 | 0.615 | 0.635 | 0.673 | 0.673 | 0.778 | 0.769 | 0.692 | 0.647 | 0.598 | 1.000 | 0.785 | 0.730 | 0.752 | 0.729 |
| क् | 0.640 | 0.640 | 0.620 | 0.640 | 0.680 | 0.680 | 0.741 | 0.724 | 0.700 | 0.655 | 0.680 | 0.824 | 1.000 | 0.840 | 0.764 | 0.834 |
| * | 0.706 | 0.706 | 0.676 | 0.706 | 0.765 | 0.765 | 0.706 | 0.651 | 0.715 | 0.824 | 0.765 | 0.730 | 0.840 | 1.000 | 0.972 | 0.947 |
| * | 0.694 | 0.694 | 0.667 | 0.694 | 0.750 | 0.750 | 0.801 | 0.698 | 0.778 | 0.659 | 0.583 | 0.752 | 0.812 | 0.915 | 1.000 | 0.920 |
| * | 0.684 | 0.684 | 0.658 | 0.684 | 0.737 | 0.737 | 0.669 | 0.682 | 0.763 | 0.718 | 0.573 | 0.729 | 0.741 | 0.947 | 0.920 | 0.947 |

Fig. 2. Pairwise edit distances computed using weighted trees.

Given this representation we can cast the problem of computing distances between different shapes as that of finding the tree edit distance between the weighted graphs for their skeletons.

With our measure assigned to each edge of the tree, we define the cost of matching two edges as the difference of the total length ratio measure along the branches. The cost of eliminating an edge is equivalent to the cost of matching it to an edge with zero weight, i.e. one along which the total length ratio is zero.

Using the edit distance of the shock trees we generate two similarity measures for a pair of shapes.

- The first measure is obtained weighting the nodes with the border length ratio normalized by the total length of the border of the shape. That is the length of the fraction of the border spanned by the shock group divided by the total length of the border. In this way the sum of the weights in a tree is 1 and the measure is scale invariant. The similarity of the shapes is computed by adding the minimum weight for each matched node, that is $d_{w,w'} = \sum_i \min(w_i, w_i')$, where w_i and w_i' are the weight of the nodes that are matched together by our tree edit distance algorithm.
- The second measure of shape similarity is computed from the unweighted structure: We assign a uniform edit cost of 1 to each node and we compute the average ratio of matched nodes: $d_{w,w'} = \frac{1}{2} \left(\frac{\#\hat{T}}{\#T_1} + \frac{\#\hat{T}}{\#T_2} \right)$, where T_1 and

 T_2 are the two trees to be matched, \hat{T} is the median of the two trees obtained through cut operations only, and # indicates the number of nodes in the tree.

3 Graph-Clustering

We pose the problem of learning the set of shape-classes as that of finding pairwise clusters in the distribution of tree-edit distance. The process of pairwise clustering is somewhat different to the more familiar one of central clustering. Whereas central clustering aims to characterise cluster-membership using the cluster mean and variance, in pairwise clustering it is the relational similarity of pairs of objects which are used to establish cluster membership. Although less well studied than central clustering, there has recently been renewed interest in pairwise clustering aimed at placing the method on a more principled footing using techniques such as mean-field annealing [14].

To commence, we require some formalism. We are interested in grouping a set of graphs $\mathcal{G}=\{G_1,....,G_{|M|}\}$ whose index set is M. The set of graphs is characterised using a matrix of pairwise similarity weights. The elements of this weight matrix are computed using tree-edit distance $d_{i,j}$ between the graphs indexed i and j. Here we use the exponential similarity function $W_{i,j}^{(0)}=\{\exp[-kd_{i,j}] \text{ if } i\neq j, 0 \text{ otherwise}\}$ to generate the elements of the weight-matrix, where k is a constant which is heuristically set. The aim in graph-clustering is to locate the updated set of similarity weights which partition the set of graphs into disjoint subsets. To be more formal, suppose that Ω is the set of graph-clusters and let S_{ω} represent the set of the graphs belonging to the cluster indexed ω . Further, let $s_{i\omega}^{(n)}$ represent the probability that the graph indexed i belongs to the cluster indexed ω at iteration i of the algorithm. We are interested in posing the clustering problem in a maximum likelihood setting. Under the assumption that the cluster memberships of the graphs follow a Bernoulli distribution with the link-weights as parameters, the likelihood-function for the weight matrix i is given by

$$P(W) = \prod_{\omega \in \Omega} \prod_{(i,j) \in M \times M} W_{i,j}^{s_{i\omega}s_{j\omega}} (1 - W_{i,j})^{1 - s_{i\omega}s_{j\omega}}$$
(1)

The corresponding log-likelihood function is

$$\mathcal{L} = \sum_{\omega \in \Omega} \sum_{(i,j) \in M \times M} \left\{ s_{i\omega} s_{j\omega} \ln W_{ij} + (1 - s_{i\omega} s_{j\omega}) \ln(1 - W_{i,j}) \right\}$$
(2)

We have recently, shown how this log-likelhood function can be iteratively opt-mised using an EM-like process. In the E (expectation) step, the cluster membership probabilities are updated according to the formula

 $s_{i\omega}^{(n+1)} = \frac{\prod_{j \in M} \left\{ \frac{W_{i,j}^{(n)}}{1 - W_{ij}^{(n)}} \right\}^{s_{j\omega}^{(n)}}}{\sum_{i \in M} \prod_{j \in M} \left\{ \frac{W_{ij}^{(n)}}{1 - W_{ij}^{(n)}} \right\}^{s_{j\omega}^{(n)}}}$ (3)

Once the revised cluster membership variables are to hand then we apply the M (maximisation) step of the algorithm to update the similarity-weight matrix. The updated similarity-weights are given by $W_{ij}^{(n+1)} = \sum_{\omega \in \Omega} s_{i\omega}^{(n)} s_{j\omega}^{(n)}$. These two steps are interleaved and iterated to convergence.

To set the number of clusters we perform a modal analysis on the initial similarity matrix $W^{(0)}$. Here we use a result due to Sarkar and Boyer [17] who have shown that the positive eigenvectors of the matrix of similarity-weights can be used to assign nodes to clusters. Using the Rayleigh-Ritz theorem, they observe that the scalar quantity $x^t W^{(0)} x$ is maximised when x is the leading eigenvector of $W^{(0)}$. Moreover, each of the subdominant eigenvectors corresponds to a disjoint perceptual cluster. They confine their attention to the same-sign positive eigenvectors (i.e. those whose corresponding eigenvalues are real and positive, and whose components are either all positive or are all negative in sign). If a component of a positive eigenvector is non-zero, then the corresponding node belongs to the perceptual cluster associated with the associated eigenmodes of the weighted adjacency matrix. The eigenvalues $\lambda_1, \lambda_2, \ldots$ of $W^{(0)}$ are the solutions of the equation $|W^{(0)} - \lambda I| = 0$ where I is the $|V| \times |V|$ identity matrix. The corresponding eigenvectors $x_{\lambda_1}, x_{\lambda_2}, \dots$ are found by solving the equation $W^{(0)}x_{\lambda_i}=\lambda_i x_{\lambda_i}$. Let the set of positive same-sign eigenvectors be represented by $\Omega = \{\omega | \lambda_{\omega} > 0 \land [(\boldsymbol{x}_{\omega}^{*}(i) > 0 \forall i) \lor \boldsymbol{x}_{\omega}^{*}(i) < 0 \forall i])\}$. Since the positive eigenvectors are orthogonal, this means that there is only one value of ω for which $x_{\omega}^{*}(i) \neq 0$. In other words, each node i is associated with a unique

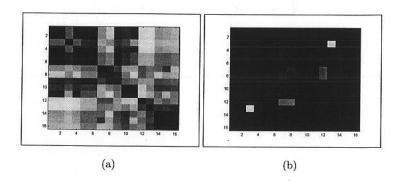


Fig. 3. (a) Initial similarity matrix for the unweighted tree edit distances; (b)Final similarity matrix for the unweighted tree edit distances.

cluster. We denote the set of nodes assigned to the cluster with modal index ω as $V_{\omega} = \{i | \boldsymbol{x}_{\omega}^{*}(i) \neq 0\}$. Hence each positive same-sign eigenvector is associated with a distinct mixing component. We use the eigenvectors of the initial affinity matrix to initialise the cluster membership variables. This is done using the magnitudes of the modal co-efficients and we set $s_{iw}^{(0)} = \frac{|x_{\omega}^*(i)|}{\sum_{i \in V_{\omega}} |x_{\omega}^*(i)|}$

4 Experiments

The 16 shapes used in our study are shown in Figures 1 and 2. In Figure 1 we show the pattern of unweighted edit distances between the shock-trees for the shapes, while Figure 2 shows the corresponding weighted tree edit distances.

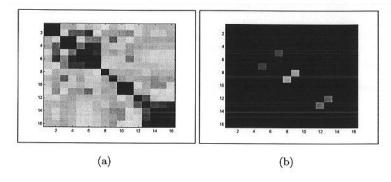


Fig. 4. (b) Initial similarity matrix for the weighted tree edit distances; (b) Final similarity matrix for the weighted tree edit distances.

In Figures 3a we show the matrix of pairwise similarity weights for the unweighted trees for the different shapes. Here the redder the entries, the stronger the similarity; the bluer the entries, the weaker the similarity. The order of the entries in the matrix is the same as the order of the shapes in Figures 1a and 1b. After six iterations of the clustering algorithm the similarity weight matrix shown in Figure 3b is obtained. There are six clusters (brush (1) + brush (2) + wrench (4); spanner (3) + horse (13); pliers (5) + pliers (6) + hammer (9) ; pliers (7) + hammer (8) + horse (12); fish (10) + fish (12); hand (14) + hand (15) + hand (16). Clearly there is merging and leakage between the different shape categories. In Figures 4a and 4b we show the initial and final similarity matrices when weighted trees are used. The entries in the initial similarity matrix are better grouped than those obtained when the unweighted tree edit distance is used. There are now seven clusters. brush (1) + brush (2); spanner (3) + spanner (4); pliers (5) + pliars (6) + pliers (7); hammer (8) + hammer (9); fish

(10) + fish (11); horse (12) + horse (13); hand (14) + hand (15) + hand (16)). These correspond exactly to the shape categories in the data-base.

Conclusions

This paper has presented a study of the problem of clustering shock-trees. We gauge the similarity of the trees using weighted and unweighted edit distance. To idetify distinct groups of trees, we develop a maximum likelihood algorithm for pairwise clustering. This takes as its input a matrix of pairwise similarities between shock-trees computed from the edit distances. The algorithm is reminiscent of the EM algorithm and has interleaved iterative steps for computing cluster-memberships and for updating the pairwise similarity matrix. The number of clusters is controlled by the number of same-sign eigenvectors of the current similarity matrix. Experimental evaluation of the method shows that it is capable of extracting clusters of trees which correspond closely to the shapecategories present.

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Feature Selection for Classification Using Genetic Algorithms with a Novel Encoding

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Abstract. Genetic algorithms with a novel encoding scheme for feature selection are introduced. The proposed genetic algorithm is restricted to a particular predetermined feature subset size where the local optimal set of features is searched for. The encoding scheme limits the length of the individual to the specified subset size, whereby each gene has a value in the range from 1 to the total number of available features.

This article also gives a comparative study of suboptimal feature selection methods using real-world data. The validation of the optimized results shows that the true feature subset size is significantly smaller than the global optimum found by the optimization algorithms.

Keywords: pattern recognition, feature selection, genetic algorithm

1 Introduction

In real-world classification problems the relevant features are often unknown a priori. Thus, many features are derived and the features which do not contribute or even worsen the classification performance have to be discarded. Therefore, many algorithms exist which typically consist of four basic steps [3]:

- 1. a generation procedure to generate the next subset of features X.
- 2. an evaluation criterion J to evaluate the quality of X.
- 3. a stopping criterion for concluding the search. It can either be based on the generation procedure or on the evaluation function.
- 4. a validation procedure for verifying the validity of the selected subset.

The task of feature selection is to reduce the number of extracted features to a set of a few significant features which optimize the classification performance. The best subset

$$X = \{x_i | i=1,\ldots,d; x_i \in Y\}$$
 (1)

is selected from the set

$$Y = \{y_i | i = 1, \dots, D\},$$
 (2)

where D is the number of extracted features and $d \leq D$ denotes the size of the feature subset [4]. A feature selection criterion function J(X) evaluates a chosen

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